Large and Moderate Deviations in Poisson Navigations

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(This talk is a joint work with B. Jahnel and S. K. Jhawar)

Ghosh

 $lackbox{ }$ Let \mathcal{P}_{λ} be a homogeneous Poisson point process on \mathbb{R}^2 with intensity $\lambda.$



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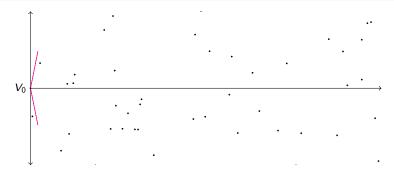
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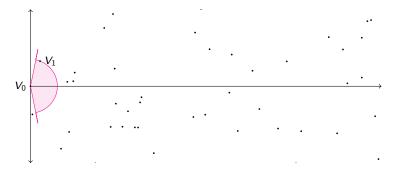
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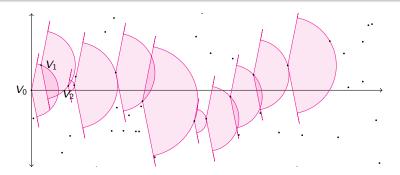


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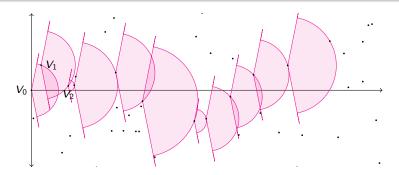


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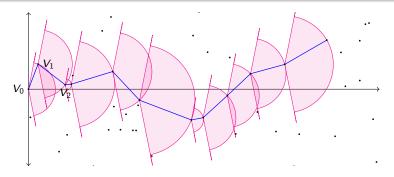
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- The node V_{i+1} is called the successor of $V_i \in \mathcal{P}_{\lambda}$ in the navigation.

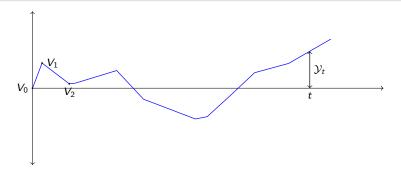


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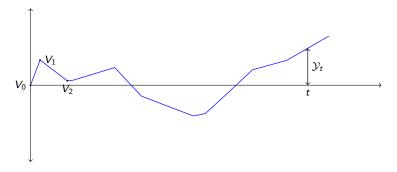
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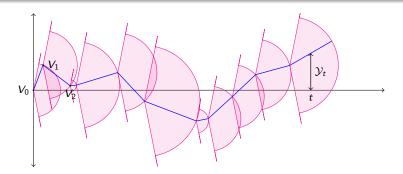


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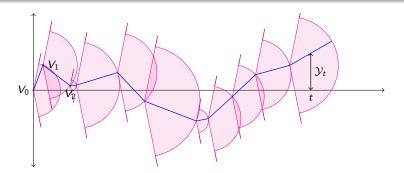
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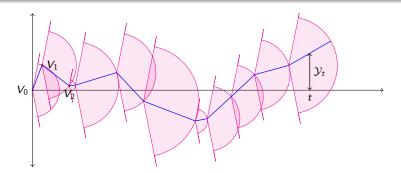
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- The Main Challenge: Steps may not be independent. (Independent for $\theta \leq \pi/4$).
- We need to account for the random horizontal step-sizes, which are not independent of the vertical displacements.



Moderate Deviation Principle

Theorem 1 (MDP) [G., Jahnel, and Jhawar (2025)]

For any $0<\lambda,\ 0<\theta\leq \frac{\pi}{2}$, and $0<\varepsilon<\frac{1}{2}$, the family $\left\{t^{-1/2-\varepsilon}\mathcal{Y}_t\right\}_{t\geq 0}$ obeys the moderate deviation principle with rate $t^{2\varepsilon}$ and rate function $I_{\lambda,\theta}(x):=\rho(\lambda,\theta)x^2$, where $\rho(\lambda,\theta)>0$. This means that for any Borel set $\Gamma\subseteq\mathbb{R}$,

$$\begin{split} -\inf_{x\in\Gamma^\circ}I_{\lambda,\theta}(x) &\leq \liminf_{t\to\infty}t^{-2\varepsilon}\log\mathbb{P}\big(t^{-1/2-\varepsilon}\mathcal{Y}_t\in\Gamma\big)\\ &\leq \limsup_{t\to\infty}t^{-2\varepsilon}\log\mathbb{P}\big(t^{-1/2-\varepsilon}\mathcal{Y}_t\in\Gamma\big) \leq -\inf_{x\in\overline{\Gamma}}I_{\lambda,\theta}(x). \end{split}$$

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$$ho(\lambda, heta) := rac{\sqrt{\pi \lambda heta} \, \sin heta}{2 heta - \sin(2 heta)}.$$

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- β controls the optimal balance between making unusually many steps in order to reach t horizontally, which might be beneficial to reach level x vertically.

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- What about $\theta > \frac{\pi}{4}$? Challenges:
 - Dependencies between steps. This can be managed using a renewal structure.
 - The renewal steps have exponential tails.

We have results assuming some control on renewal steps.

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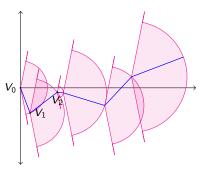
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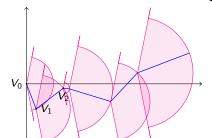
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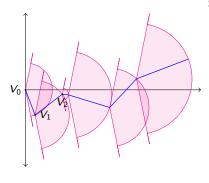
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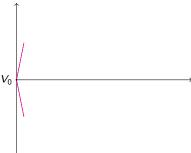
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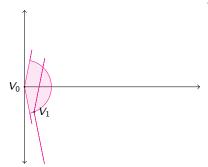
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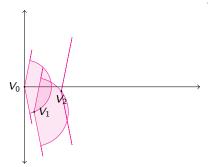
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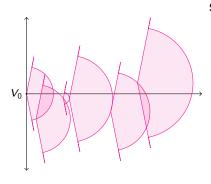
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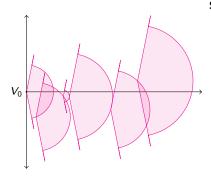
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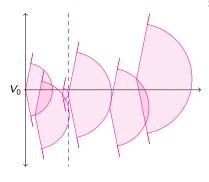
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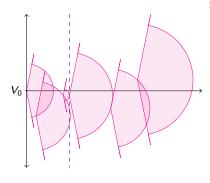
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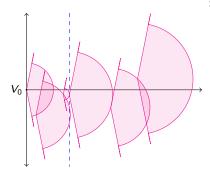
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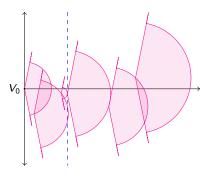


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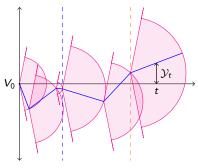
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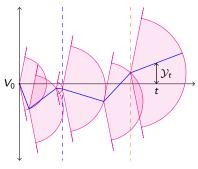
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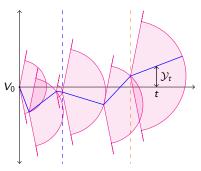
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- Define "Renewal Steps". $\tau_1 := \inf\{n>0: H_n=\emptyset\}, \\ \tau_k := \inf\{n>\tau_{k-1}: H_n=\emptyset\}.$ Segments between two consecutive stopping times are independent.
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Step 1: Renewal Structure.

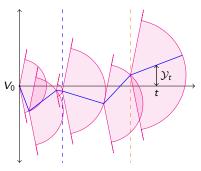
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- Then

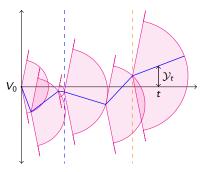
$$\boxed{ \mathcal{Y}_t = \sum_{i=1}^{\mathcal{K}_t'} Y_i' + \sum_{i=\tau_{\mathcal{K}_t'}+1}^{\mathcal{K}_t} Y_i + (\mathsf{Rest \; step}), } \quad \text{ where } \quad Y_i' = \sum_{j=\tau_{i-1}+1}^{\tau_i} Y_j. }$$



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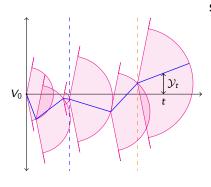
$$\boxed{ \mathcal{Y}_t = \sum_{i=1}^{K_t'} Y_i' + \sum_{i=\tau_{K_t'}+1}^{K_t} Y_i + (\mathsf{Rest \; step}), } \quad \text{ where } \quad Y_i' = \sum_{j=\tau_{i-1}+1}^{\tau_i} Y_j. }$$



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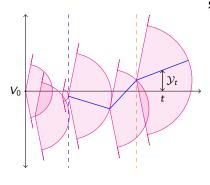
$$\mathcal{Y}_t = \mathcal{Y}_t' + \sum_{i=\tau_{K_t'}+1}^{K_t} Y_i + \text{(Rest step)}, \qquad \text{where} \qquad Y_i' = \sum_{j=\tau_{i-1}+1}^{\tau_i} Y_j.$$



Step 2: Exponential Equivalence.

ullet \mathcal{Y}_t and \mathcal{Y}_t' are exponentially equivalent: $orall \delta > 0$,

$$\limsup_{t\uparrow\infty} t^{-2\epsilon}\,\log \mathbb{P}\big(|\mathcal{Y}_t-\mathcal{Y}_t'|\geq \delta\,t^{1/2+\epsilon}\big)=-\infty.$$



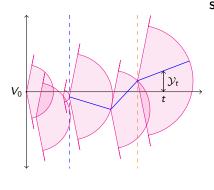
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$$\bullet \ \mathbb{P}(|\mathcal{Y}_t - \mathcal{Y}_t'| \ge \delta t^{1/2 + \epsilon})$$

$$\le \mathbb{P}(\sum_{i=\tau_{K_t'}+1}^{K_t+1} |Y_i| \ge \delta t^{1/2 + \epsilon})$$

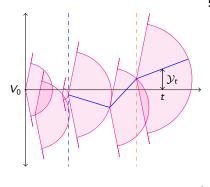


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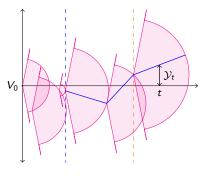
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It is unlikely to make more that t^2 steps to reach t.



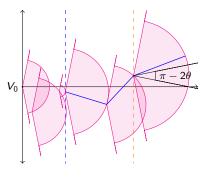
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- It is unlikely to make more that t² steps to reach t.
- τ_1 has exponential tail. Involves Markov chain argument (Most technical part).



Step 2: Exponential Equivalence.

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$$\mathbb{P}(|\mathcal{Y}_t - \mathcal{Y}_t'| \ge \delta t^{1/2+\epsilon})$$

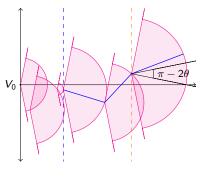
$$\le \mathbb{P}(\sum_{i=\tau_{K_t'}+1}^{K_t+1} |Y_i| \ge \delta t^{1/2+\epsilon})$$

$$\le \mathbb{P}(K_t' \ge t^2) + \mathbb{P}(\tau_1 > t^{\epsilon'})$$

$$+\mathbb{P}ig(\sum_{i=1}^{\lfloor t^{\epsilon'} \rfloor} \overline{R}_i \geq \delta t^{1/2+\varepsilon}ig)$$

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 Let L_n be the width of H_n.

$$L_n \leq \stackrel{\cdot \cdot \cdot}{M_n} := (M_{n-1} - \lfloor \frac{n}{R_n} \cos \theta \rfloor) + 1\{\underline{\Phi}_n \in T_\theta\} + \max\{M_{n-1}, \lceil \overline{R}_n \rceil\} 1\{\underline{\Phi}_n \notin T_\theta\}.$$



Step 2: Exponential Equivalence.

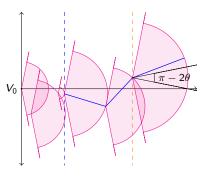
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Ghosh



Step 2: Exponential Equivalence.

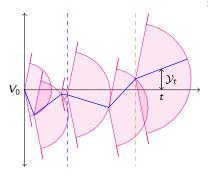
• \mathcal{Y}_t and \mathcal{Y}_t' are exponentially equivalent: $\forall \delta > 0$,

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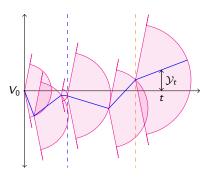


Step 3: No Randomness in Horizontal Direction.

ullet \mathcal{Y}_t' and $\mathcal{Y}_t^{\varkappa}$ are exponentially equivalent: $\forall \delta > 0$,

$$\limsup_{t \uparrow \infty} t^{-2\epsilon} \log \mathbb{P}(|\mathcal{Y}_t' - \mathcal{Y}_t^{\varkappa}| \ge \delta t^{1/2 + \epsilon}) = -\infty,$$

where
$$\varkappa := \mathbb{E}[X_1']$$
 and $\mathcal{Y}_t^{\varkappa} = \sum_{i=1}^{\lfloor t/\varkappa \rfloor} Y_i'$.



Step 3: No Randomness in Horizontal Direction.

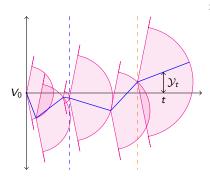
• \mathcal{Y}'_t and \mathcal{Y}''_t are exponentially equivalent: $\forall \delta > 0$,

$$\limsup_{t\uparrow\infty} t^{-2\epsilon} \, \log \mathbb{P}\big(|\mathcal{Y}_t' - \mathcal{Y}_t^\varkappa| \geq \delta \, t^{1/2+\epsilon}\big) = -\infty,$$

where
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 and $\mathcal{Y}_t^{\varkappa} = \sum_{i=1}^{\lfloor t/\varkappa \rfloor} Y_i'$.

Remarks:

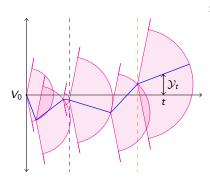
- Infact, the the process makes the typical number of steps.
- Crucially uses exponential unlikeliness of making unexpected many steps.



Step 4: MDP for $\mathcal{Y}_t^{\varkappa}$.

• $\left\{t^{-1/2-\varepsilon}\mathcal{Y}_t^{\varkappa}\right\}_{t\geq 0}$ obeys the moderate-deviation principle with rate $t^{2\varepsilon}$ and rate function

$$I_{\lambda,\theta}^{\varkappa}(x) := \frac{\varkappa \cdot x^2}{2 \cdot \mathbb{E}[Y_1'^2]} = \frac{\mathbb{E}[X_1'] \cdot x^2}{2 \cdot \mathbb{E}[Y_1'^2]}.$$

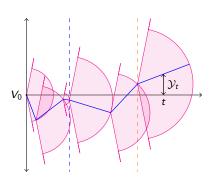


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• Proof: Application of Eichelsbacher–Löwe criteria. [ESAIM Probab. Stat., 7:209–218.]



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(proved)

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Thank You